

# SVARs & identification + Intro to continuous time stochastic processes

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# Outline of today's talk

- SVARs.
- Identification methods.
- Continuous time stochastic processes.

# Reading on SVARs and identification

- Canova: “Methods for applied macroeconomic research”.
  - Section 4.5 & 4.6 covers Identification and SVARs.
  - Section 10.3 covers this in a Bayesian context.
- Wikipedia, as needed for basic results in linear algebra.
  - Reading all of the pages in this category would be a good start:  
[https://en.wikipedia.org/wiki/Category:Matrix\\_decompositions](https://en.wikipedia.org/wiki/Category:Matrix_decompositions)
- Christiano, Eichenbaum and Evans (2005):
  - A classic paper which you ought to be familiar with.
  - [www.tau.ac.il/~yashiv/cee.pdf](http://www.tau.ac.il/~yashiv/cee.pdf)

# Readings on continuous time processes etc.

- Cochrane (2012):
  - Nice review of continuous time stochastic processes, with a macro slant.
  - [http://faculty.chicagobooth.edu/john.cochrane/research/papers/continuous\\_time\\_linear\\_models.pdf](http://faculty.chicagobooth.edu/john.cochrane/research/papers/continuous_time_linear_models.pdf)
- Any finance textbook for details (shouldn't really be needed).
  - E.g. Chapter 3 of Merton's "Continuous Time Finance".
- Wikipedia as needed...

# Structural VARs: Motivation (1/2)

- We would like to know what the effects of (say) an unexpected increase in monetary policy is.
- But a change in monetary policy will produce changes in other variables within the same time period.
- Conversely, exogenous shocks to other variables will produce automatic reactions from monetary policy.
  - E.g. a Taylor Rule.
- Thus, if we see that interest rates were (say) tighter than was expected yesterday, we do not know if this was due to a change in policy or if it was an endogenous reaction to other changes in the economy.
  - A standard VAR tells us nothing about the effects of changes in policy!

## Structural VARs: Motivation (2/2)

- Furthermore, even after contemporaneous responses of one variable to another have been taken into account, there may still be correlations in the shocks.
  - For example, an exogenous increase in rainfall may both decrease labour supply holding fixed the wage, and increase labour demand holding fixed the wage.
  - Thus in a VAR in which rainfall is omitted, it would show up as both a labour supply and a labour demand shock.
  - No way of knowing how much of this variance component due to rainfall should be assigned to supply, and how much should be assigned to demand.
  - However, this really reflects a failure of the model (omitting an observable variable).
- Alternatively, some variables may respond directly to structural shocks to other variables.
- In macroeconomic terms this is rather implausible, as shocks are generally not observed directly, and if they are observed, they're generally only observed by the agent that experiences the shock.
- Nonetheless, in a few rare cases this may be justified.

# Structural VARs: Definition

- This suggests the following structural representation:

$$x_t = c + a_0 x_t + a_1 x_{t-1} + \dots + a_p x_{t-p} + u_t + b_0 u_t$$

- where both  $a_0$  and  $b_0$  have a zero diagonal and where  $u_t \sim \text{WNIID}(0, \Sigma_u)$ , with  $\Sigma_u$  diagonal.

- Then:

$$(I - a_0)x_t = c + a_1 x_{t-1} + \dots + a_p x_{t-p} + (I + b_0)u_t$$

- Then if we define  $A := I - a_0$  and  $B := I + b_0$ :

$$x_t = A^{-1}c + A^{-1}a_1 x_{t-1} + \dots + A^{-1}a_p x_{t-p} + A^{-1}Bu_t.$$

- Compare this to our previous reduced form:

$$x_t = \mu + \phi_1 x_{t-1} + \phi_2 x_{t-2} + \dots + \phi_p x_{t-p} + \varepsilon_t, \quad \varepsilon_t \sim \text{WNIID}(0, \Sigma_\varepsilon)$$

- Matching terms gives:

$$A\mu = c, \quad A\phi_1 = a_1, \quad \dots, \quad A\phi_p = a_p, \quad A\varepsilon_t = Bu_t, \\ A\Sigma_\varepsilon A' = B\Sigma_u B'$$

# Structural VARs: Basic identification (1/2)

- Our hope is to be able to use some prior restrictions on  $A$  and  $B$  (or equivalently  $a_0$  and  $b_0$ ), in order to solve for  $\Sigma_u$  in the equation  $A\Sigma_\varepsilon A' = B\Sigma_u B'$ .
- We know  $A$  and  $B$  have a unit diagonal, and that  $\Sigma_u$  is zero everywhere except the diagonal.
- If we knew  $A\Sigma_\varepsilon A'$  could we at least work out  $B$  and  $\Sigma_u$ ?
  - No, not uniquely, without additional information.
  - By the Cholesky decomposition, there exists a lower triangular matrix  $L$  such that  $A\Sigma_\varepsilon A' = LL'$ . So one candidate solution is  $B := L(\text{diag diag } L)^{-1}$ , and  $\Sigma_u := (\text{diag diag } L)^2$ .
  - But let  $U$  be any real orthogonal matrix. Then  $A\Sigma_\varepsilon A' = (LU)(LU)'$  too. Thus  $B := LU(\text{diag diag } LU)^{-1}$ , and  $\Sigma_u := (\text{diag diag } LU)^2$  is another solution.
  - The space of all  $n \times n$  orthogonal matrices is  $\frac{n(n-1)}{2}$  dimensional, so this is the number of restriction we need on  $B$  if we already know  $A\Sigma_\varepsilon A'$ .
  - This may be seen directly from noting that in the equation  $RR' = S$ , with  $S$  symmetric, the equations above the diagonal are identical to those below.



# Structural VARs: Basic identification (2/2)

- In practice  $B$  is almost always assumed to be equal to the identity matrix, for the reasons I gave previously.
  - If it's not, it reflects either strange informational assumptions, or omitted variables.
- So with  $B$  known, can we pin down  $A$  without additional assumptions?
  - No. Much as before, the equation  $A\Sigma_\varepsilon A' = B\Sigma_u B'$  has  $\frac{n(n-1)}{2}$  free parameters with  $B$  known.
  - So this is the number of assumptions we need to make on  $A$ .
- A common assumption is that  $A$  is lower triangular, which gives the required  $\frac{n(n-1)}{2}$  restrictions.
  - This means  $a_0$  is strictly lower triangular, implying a “causal ordering” on the variables.
  - The variable ordered first is assumed to have no contemporaneous response to later variables.
  - The one ordered second just responds to the first contemporaneously, but no others. Etc. etc. till...
  - The one ordered last responds contemporaneously to all variables.

# Reporting SVAR results: Impulse responses

- Normally, we are primarily interested in the response of the economy to some shock.
  - So, suppose there's a shock of one standard deviation to variable  $i$  in period 0, and then from then on, no further shocks hit the economy.
  - I.e., suppose that  $u_{0,i} = 1$  for some  $i \in \{1, \dots, n\}$ , but that  $u_{t,j} = 0$  for all  $t \in \mathbb{N}$  and all  $j \in \{1, \dots, n\}$ , unless  $t = 0$  and  $j = i$ .
  - Given these assumptions on the shocks, we can simulate the SVAR model and plot the responses of the variables of interest.
  - The result of this is an “impulse response”. Often plotted relative to the variable's mean.
- In non-linear models, there are various possible definitions of an impulse response.
  - One is  $\mathbb{E}[x_t | u_{0,i} = 1]$ .
  - Another is  $\mathbb{E}[x_t | u_{0,i} = \tilde{u}_{0,i} + 1]$ , where  $\tilde{u}_{0,i}$  has the same distribution as that of  $u_{t,i}$ .
    - Some authors also condition on the initial state in these expectations. (Dynare does not.)

# Other identification methods: AB restrictions

- Causal orderings are deeply implausible. Most variables have some contemporaneous effect on most other variables.
- Indeed, many variables have strong anticipatory effects on other variables.
  - If a shock to another variable is expected in future (and the econometricians dataset is insufficient to pick this up) then shocks (observed) tomorrow might have an effect on variables today.
- Other restrictions on the  $A$  and  $B$  matrices based on theory are often as bad, for basically the same reason.
- The Blanchard and Perotti (2002) approach uses micro data to estimate some parameters of the  $A$  matrix in a fiscal policy context.
  - Their argument is based upon government taking more than quarter to respond, so is little better than the causal ordering approach.

# Other identification methods: Sign restrictions

- The sign restriction approach (Uhlig 2005) effectively places a flat prior over the space of all orthogonal rotation matrices, then truncates this prior to zero in areas where the model generates “the wrong results” in some sense.
  - “Flat” may not be as intuitive as it seems, see e.g. Baumeister & Hamilton (2017).
- “Wrong” is usually defined in terms of the sign of the impulse response to a certain shock at a certain point in time.
  - May end up assuming what it wants to prove. E.g. causal ordering identification of monetary policy shocks often produces “price puzzles”, with increasing interest rates increasing inflation.
  - Assuming away price puzzles begs the question of whether these are real features or not.
- If interpreted classically, sign restrictions only produce set identification, not point identification. (see Moon, Schorfheide and Granziera 2013).
  - Following identification via sign restrictions, there is no such thing as “the” estimated impulse response.
  - Rather, the estimator provides a band of impulse responses, even with infinite data.
  - Finite sample parameter uncertainty produces even larger bands.

# Other identification methods: Narrative evidence

- Pioneered by Romer and Romer (1989), who use the text of FOMC meetings to identify times when policy makers intended to use contractionary policy to bring down inflation.
  - Later work has tightened the definition of a monetary shock.
- In fiscal contexts, Ramey and Shapiro (1998) performed a similar analysis using military build-ups.
- One difficulty with this approach is that hand selected shocks will always “smell funny”.
  - Some recent researchers have ameliorated this via using automated textual analysis.
- Another problem is that it’s not always clear that the narrative analysis procedure really succeeds in purging all endogeneity.

# Other identification methods: Long-run restrictions

- While there's a lot of debate about how the economy evolves in the short-run, there's a lot more consensus about the long-run effects of various shocks.
  - E.g. only a technology shock increases GDP per capita in the long-run. Monetary shocks are neutral for all variables in the long-run. Etc.
- Blanchard and Quah (1989) exploit this for identification.
  - It is a bit like a sign restriction at  $t = \infty$ , but since they are imposing exact coefficients for the long-run response they get point, not set, identification.
  - Can be imposed without simulation, either:
    - by deriving the limit of the IRF by diagonalization (see exercises), or,
    - by using the following Beveridge-Nelson type decomposition:  $\Delta x_t = \Theta(L)\varepsilon_t = \Theta(1)\varepsilon_t + \frac{\Theta(L) - \Theta(1)}{1-L}\Delta\varepsilon_t$ , so the permanent impact of a structural shock is  $\Theta(1)A^{-1}Bu_t$ .

# SVARs and identification in practice

- There is a huge literature looking at the responses of monetary and fiscal shocks.
- Results vary wildly depending on which identification method is used, though there is more consensus about monetary shocks than fiscal ones.
  - For example, in a cross country study Iletzki, Mendoza and Vegh (2013) find basically zero fiscal multipliers in developed, open economies, and Ramey's narrative based work finds at most moderate multipliers, around one.
  - On the other hand Perotti continues to find large multipliers.
- The correct response is broad distrust of most VAR identification methods.
- In any case, it is unclear why we should care about fiscal multipliers.
  - The fact that government expenditure increases GDP more than one for one tells us nothing about whether this is good for welfare.
  - In fact, in most modern macro models that generate large multipliers, expansionary fiscal policy is unambiguously bad for welfare.

# Continuous time stochastic processes

- The stochastic processes we looked at in the first lecture were random variables taking their value from the vector space of sequences (i.e. functions  $\mathbb{Z} \rightarrow \mathbb{R}$ ).
- In some circumstances, it is easier to work in continuous time, i.e. with random variables taking their value from the vector space of functions  $\mathbb{R} \rightarrow \mathbb{R}$ .
  - This is the standard in finance.
  - It's also increasingly more common in macro, and we'll look at some continuous time DSGE models later in the course.



# The Wiener process

- The Wiener process (aka “standard Brownian motion”) is the building blocks of most continuous time stochastic processes.
- You might find it helpful to think of the Wiener process as the continuous time analogue of a random walk.
- The process,  $W_t$  is characterised by the following properties:
  1.  $W_0 = 0$ .
  2.  $W_t$  is almost surely everywhere continuous.
  3. If  $s_1, s_2, \dots, s_{n+1}, t_1, t_2, \dots, t_n \in \mathbb{R}^+$  satisfy  $0 < s_i < t_i \leq s_{i+1}$  for all  $i \in \{1, \dots, n\}$ , then  $W_{t_i} - W_{s_i}$  is independent of  $W_{t_j} - W_{s_j}$  for all  $i \neq j$ .
  4. If  $0 \leq s < t$ , then  $W_t - W_s \sim N(0, t - s)$ .
- The process  $\mu t + \sigma W_t$  is called a Wiener process with drift  $\mu$  and infinitesimal variance  $\sigma^2$ .

# The Itô integral

- We would often like to work with processes with time varying drift and time varying infinitesimal variance.
  - Scaling the Wiener process by a time varying amount will not work, as this will change both the level of the process and its future infinitesimal variance.
  - In some loose sense then, we need to “differentiate” the process, scale it, and then integrate back.
  - However, the Wiener process is not differentiable.
- Itô defined a new integral (with different integration laws) in order to tackle this.
  - It allows us to integrate a function times a kind of “derivative” of the Wiener process.
- In particular, if:
  - for all  $n \in \mathbb{N}$ ,  $\pi_n$  is an increasing sequence of length  $n + 1$ , with  $\pi_{n,0} = S$  and  $\pi_{n,n} = T$ , and  $\lim_{n \rightarrow \infty} \max_{i \in \{1, \dots, n\}} |\pi_{n,i} - \pi_{n,i-1}| = 0$ ,
  - $W_t$  is a Wiener process, and  $X_t$  is another continuous time stochastic process that is left-continuous and locally bounded,
  - then we define:

$$\int_S^T X_t dW_t := \text{plim}_{n \rightarrow \infty} \sum_{i \in \{1, \dots, n\}} X_{\pi_{n,i-1}} (W_{\pi_{n,i}} - W_{\pi_{n,i-1}}).$$

# Useful property: The Itô isometry

- Given  $S_X < T_X, S_Y < T_Y$ , with  $\max\{S_X, S_Y\} < \min\{T_X, T_Y\}$  and given continuous stochastic processes  $X_t$  and  $Y_t$ :

$$\mathbb{E} \left[ \left( \int_{S_X}^{T_X} X_t dW_t \right) \left( \int_{S_Y}^{T_Y} Y_t dW_t \right) \right] = \mathbb{E} \left[ \int_{\max\{S_X, S_Y\}}^{\min\{T_X, T_Y\}} X_t Y_t dt \right]$$

- Informal “proof”:

- Note:

$$\mathbb{E} \left[ \left( \int_{S_X}^{T_X} X_t dW_t \right) \left( \int_{S_Y}^{T_Y} Y_t dW_t \right) \right] = \mathbb{E} \left[ \left( \int_{S_X}^{T_X} \int_{S_Y}^{T_Y} X_s Y_t dW_s dW_t \right) \right]$$

- $\mathbb{E} dW_s dW_t$  “equals”  $0 dt$  if  $s \neq t$  and “equals”  $1 dt$  if  $s = t$ .

# Drift diffusion processes

- Processes used in finance (and continuous time macro) often take the form:

$$X_t = X_0 + \int_0^t \mu(X_u, u) du + \int_0^t \sigma(X_u, u) dW_u .$$

- The first integral here is a standard integral, the second is an Itô one!
  - The function  $\mu(X_t, t)$  controls the drift of the process at  $t$ .
  - The function  $\sigma(X_t, t)$  controls the infinitesimal variance at  $t$ .
- 
- In practice, this expression is usually written in the more compact “stochastic differential equation” form:

$$dX_t = \mu(X_t, t) dt + \sigma(X_t, t) dW_t .$$

- However, it is important to remember that the latter expression is just a shorthand for the former.

# Itô's lemma

- Itô's lemma is an equivalent of the chain rule for continuous time stochastic processes.

- Suppose:

$$dX_t = \mu_t dt + \sigma_t dW_t.$$

- Then for any twice differentiable function  $f: \mathbb{R}^2 \rightarrow \mathbb{R}$ , Itô's lemma states:

$$df(t, X_t) = \left( f_1 + \mu_t f_2 + \frac{1}{2} \sigma_t^2 f_{22} \right) dt + \sigma_t f_2 dW_t$$

- For example, let  $Y_t = \exp X_t$ , then:

$$dY_t = d \exp X_t = \left( \mu_t Y_t + \frac{1}{2} \sigma_t^2 Y_t \right) dt + \sigma_t Y_t dW_t.$$

# Ornstein-Uhlenbeck processes

- Ornstein-Uhlenbeck processes are the continuous time equivalent of AR(1) processes.
  - Recall for later that the AR(1) process  $x_t = (1 - \rho)\mu + \rho x_{t-1} + \sigma \varepsilon_t$  has an MA( $\infty$ ) representation  $x_t = \mu + \sigma \sum_{s=0}^{\infty} \rho^s \varepsilon_{t-s}$ .

- Ornstein-Uhlenbeck processes are solutions to the s.d.e.:

$$dX_t = \theta(\mu - X_t) dt + \sigma dW_t.$$

- To find their properties, first define  $Y_t = e^{\theta t} X_t$ , then, by Itô's lemma:

$$\begin{aligned} dY_t &= d(e^{\theta t} X_t) = (\theta Y_t + \theta(\mu - X_t)e^{\theta t}) dt + \sigma e^{\theta t} dW_t \\ &= \theta \mu e^{\theta t} dt + \sigma e^{\theta t} dW_t. \end{aligned}$$

- I.e.:

$$Y_t = Y_0 + \int_0^t \theta \mu e^{\theta u} du + \int_0^t \sigma e^{\theta u} dW_u = (X_0 - \mu) + \mu e^{\theta t} + \sigma \int_0^t e^{\theta u} dW_u.$$

- Thus:

$$X_t = (X_0 - \mu)e^{-\theta t} + \mu + \sigma \int_0^t e^{\theta(u-t)} dW_u = (X_0 - \mu)e^{-\theta t} + \mu + \sigma \int_{s=0}^t e^{-\theta s} dW_{t-s}.$$

- Define  $Z_t = X_{t+\tau}$  (i.e.  $Z_t$  is an Ornstein-Uhlenbeck process started at time  $-\tau$ ). Then, in the limit as  $\tau \rightarrow \infty$ :

$$Z_t = \mu + \sigma \int_{s=0}^{\infty} e^{-\theta s} dW_{t-s}.$$

# Conclusion and recap

- Reduced form VARs do not identify shocks.
- Identification is impossible without making strong prior assumptions.
- Continuous time stochastic processes are not so different to discrete time ones.